Genetic and environmental influences on cannabis use initiation and problematic use: a meta-analysis of twin studies

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ABSTRACT

Background Because cannabis use is associated with social, physical and psychological problems, it is important to know what causes some individuals to initiate cannabis use and a subset of those to become problematic users. Previous twin studies found evidence for both genetic and environmental influences on vulnerability, but due to considerable variation in the results it is difficult to draw clear conclusions regarding the relative magnitude of these influences.

Methods A systematic literature search identified 28 twin studies on cannabis use initiation and 24 studies on problematic cannabis use. The proportion of total variance accounted for by genes (A), shared environment (C) and unshared environment (E) in (i) initiation of cannabis use and (ii) problematic cannabis use was calculated by averaging corresponding A, C and E estimates across studies from independent cohorts and weighting by sample size.

Results For cannabis use initiation, A, C and E estimates were 48%, 25% and 27% in males and 40%, 39% and 21% in females. For problematic cannabis use A, C and E estimates were 51%, 20% and 29% for males and 59%, 15% and 26% for females. Confidence intervals of these estimates are considerably narrower than those in the source studies.

Conclusions Our results indicate that vulnerability to both cannabis use initiation and problematic use was influenced significantly by A, C and E. There was a trend for a greater C and lesser A component for cannabis use initiation compared to problematic use for females.

Keywords Cannabis, genetics, heritability, meta-analysis, twin research.

INTRODUCTION

With about 166 million annual consumers (equivalent to 3.9% of the global population aged 15–64 years), cannabis is the most widely consumed illicit drug world-wide, and by far the illicit drug consumed most commonly by young people [1]. Furthermore, a broad estimation suggests that in Europe about 1% of people consume cannabis almost daily, and several European countries have reported an increase in the number of regular or intensive users [2].

Cannabis use can lead to social harms, including accidents, violence and suicide attempts [3] and regular cannabis use can lead to physical or psychological problems, and has been found to interfere with family, school and work [4–8]. Law enforcement, public health costs and loss of productivity and work potential due to health problems are also an economic drain on society [2,5]. According to the United Nations Office on Drugs and Crime [1], cannabis use is among the most common primary reasons for entering drug-related treatment. Furthermore, cannabis use often precedes the use of other drugs, which suggests that cannabis may cause further problems as a gateway drug [9–11]. However, the exact nature of the association between cannabis use and subsequent other illicit drug use is unclear [12–14].

To deal with the problems associated with cannabis use, it is important to understand what causes some
individuals to initiate cannabis use and what causes some of those individuals to become regular users or become dependent on it. Although there may be some completely random events that cause people to vary in their cannabis use (such as changes in availability of the drug), much of the variability is likely to be due both to the nature of the environment in which they live and developed, and to their genetic make-up.

It has long been recognized that risk of cannabis and other substance (ab)use runs in families. Studies aiming to understand the basis of familial risk include family studies, adoption studies and twin studies. Family studies into cannabis use have shown moderate parent–offspring correlations (ranging between 0.30 and 0.59 [15–18]) as well as sibling–sibling correlations (ranging between 0.39 and 0.59 [15,19]). In a recent study, Merikangas et al. [20] found elevated risks for cannabis use disorders among siblings [odds ratio (OR) = 3.6], offspring (OR = 6.9) and spouses (OR = 4.4) of probands with cannabis use disorders. However, family studies cannot determine whether familial resemblance is due to genetic factors or environmental factors shared between family members. Adoption studies can distinguish genetic and shared environmental factors by comparing the similarity of the adopted child with both its adopted parents and its biological parents. To our knowledge, no adoption study has examined cannabis use specifically, but adoption studies into drug and alcohol use have found that abuse or dependence of adoptees is more related to abuse or dependence of their biological parents than their adoptive parents [21–26], indicating an important role for genetic factors.

Twin studies disentangle familial resemblance into genetic and shared environmental factors by comparing the similarity of identical (monozygotic; MZ) and non-identical ( dizygotic; DZ) twins. There have been numerous twin studies into cannabis use, but due to considerable variation in the results it is difficult to draw clear conclusions regarding the relative magnitude of genetic and environmental influences. Estimates of the proportion of variance in cannabis use accounted for by genetic influences (i.e. heritability) range from close to zero (e.g. [27,28]) to more than 60% (e.g. [29–31]). Similarly, estimates of the proportion of variance accounted for by shared environmental factors range from zero (e.g. [30,32]) to 68% [28]. Inconsistent results have also been found for problematic cannabis use (symptoms of abuse and dependence). Various explanations could be proposed for these inconsistent results, including differences in measurement scales, sample size and demographic differences (age, sex, nationality, socio-economic status). In particular, very large sample sizes are required to estimate accurately genetic and shared environmental influences when using dichotomous variables (which is the case in most cannabis use studies). For this reason, many of the individual studies barely had the power to statistically distinguish between genetic and shared environmental influences, and confidence intervals (CIs) around the estimates were often very wide.

Here we carried out a meta-analysis of existing twin studies in order to provide a more accurate estimate of the magnitude of genetic and environmental influences on cannabis use initiation and problematic cannabis use ([symptoms of] abuse/dependence). Because cannabis use in general is more prevalent among males than females [2], and some twin studies reported sex differences in contributions of genetic and environmental factors (e.g. [33,34]), meta-analyses were carried out separately for males and females in order to check for sex differences in cannabis use aetiology.

**METHOD**

**Background information—twin studies and cannabis phenotypes**

The twin design

The studies we examined in this meta-analysis are twin studies that apply genetic modelling to determine the sources of individual differences in cannabis use. Below, a short introduction to the classical twin design is provided; further details can be found elsewhere (e.g. [35–37]).

With the classical twin design, trait variance can be partitioned into its genetic and environmental (shared within twin pairs and non-shared) components, by analysing the resemblance in MZ and DZ twin pairs. Additive genetic variance (A) results from the sum of allelic effects within and across multiple genes affecting a trait. Shared environmental variance (C) is due to environmental influences shared within twin pairs, such as the family environment, prenatal influences, parental style and socio-economic status. Unshared environmental variance (E) results from environmental factors that are not shared within twin pairs (e.g. idiosyncratic events and experiences, unshared peers) and includes measurement error.

Estimates of these genetic and environmental variance components can be obtained because A, C and E each predict different patterns of MZ and DZ twin pair correlations. MZ twins share all their genes, while DZ twins share on average 50% of their genes. Hence, if A were the sole source of variance in a trait, twin correlations of 1.0 for MZ pairs and 0.5 for DZ pairs are expected. If C were the sole source of variance in a trait, a twin correlation of 1.0 for both MZ and DZ pairs is expected, and if E would be the sole source of variance in a trait a twin correlation of 0.0 for both MZ and DZ pairs is expected.

In reality, individual differences in complex phenotypes result from a combination of these genetic and environmental influences. Using the observed MZ and DZ twin pair correlations it is possible to estimate standardized A, C and E variance components, which represent the proportion of total variance accounted for by additive genetic, and shared and unshared environmental influences. All A, C and E estimates reported in this paper refer to standardized variance components. These estimates are obtained by employing maximum-likelihood modelling procedures, which determine the combination of genetic and environmental parameters that best fits the covariance structure of the observed data. In addition, confidence intervals around these estimates can also be calculated. Most reports used in our analyses employed maximum-likelihood modelling procedures using the statistical package MX [38], and others used LISREL [39,40].

It is assumed that the shared environmental variance component estimated in twin studies is generalizable to the general population. Studies including twins and their siblings make it possible to distinguish between general shared environmental influences and a special twin environment. Studies including siblings have not identified a significant twin environment effect for cannabis use [31,34,41].

Most cannabis use phenotypes are measured as dichotomous variables (i.e. cannabis users versus non-users) which can be analysed by using a threshold model [42]. This model assumes that there is an underlying continuum of liability which is distributed normally in the population. Upon this normal distribution, a threshold delimits affected versus unaffected cases. The variation in liability can be analysed in the same way as the variance for continuous variables.

**Phenotypes: initiation of cannabis use and problematic use**

The various twin studies into cannabis use have used different phenotypes (observable characteristics, traits or behaviours) such as initiation, use in the last year, regular use, symptoms of abuse or dependence to full diagnosis of abuse or dependence. In this meta-analysis we examine two cannabis-related phenotypes: initiation of cannabis use and problematic cannabis use.

Initiation of cannabis use is also often referred to as life-time cannabis use or ‘ever used cannabis’. The core aspect is that it makes a distinction between individuals who have tried cannabis at least once in their life-time versus those who have not. Hence, this phenotype is a dichotomous variable.

The other phenotype we examine, problematic cannabis use, is defined less consistently. Different definitions of problematic cannabis use can be found in the literature, ranging from symptoms of abuse to a full dependence diagnosis. Most studies use abuse and dependence criteria according to the *Diagnostic and Statistical Manual of Mental Disorders* (DSM-IV-TR [8]). According to the DSM-IV-TR ([8], p. 197) substance abuse is characterized by ‘a maladaptive pattern of substance use manifested by recurrent and significant adverse consequences related to the repeated use of substances’. Substance dependence is a more advanced state of drug abuse, the essential feature of which is ‘a cluster of cognitive, behavioural and physiological symptoms indicating that the individual continues use of the substance despite significant substance-related problems. There is a pattern of repeated self-administration that can result in tolerance, withdrawal and compulsive drug-taking behaviour’ ([8], p. 199). Withdrawal symptoms do not have to be met for a cannabis dependence diagnosis [8], although this has been subject to debate [43].

In the present analysis, problematic use is defined operationally as having one or more of the symptoms of life-time abuse or dependence. We did not limit our focus to studies that use full abuse or dependence diagnosis only, because we are interested in vulnerability to problematic cannabis use or addiction. Because of this broad definition we incorporated studies using phenotypes such as ‘abuse or dependence’, ‘one or more abuse or dependence symptoms’ and ‘abuse’. All studies incorporated into the meta-analysis have analysed problematic use as a dichotomous measure.

**Data collection: literature search and study inclusion criteria**

Selection of relevant twin studies on cannabis use and problematic use for this study started with a search of the electronic databases PubMed (http://www.ncbi.nlm.nih.gov/entrez) and ISI web of Knowledge (http://apps.isiknowledge.com) using the following keywords: *heritability/heritability/twin and cannabis/marijuana/hashish*. No restrictions regarding date range were specified. Abstracts of these search results (n = 122) were examined and relevant articles were retrieved for review. Three additional studies identified from reference lists from these studies and one manuscript in press ([44], results obtained by personal communication) were also added.

Subsequently, unsuitable studies were excluded from the analysis based on two main criteria. First, only studies specifically examining cannabis use were included, and those examining related phenotypes such as general drug use were omitted. Secondly, only studies that used twin samples and applied genetic modelling to investigate the genetics of cannabis involvement were included. This procedure identified 28 twin studies on life-time cannabis
use and 24 studies on problematic cannabis use. For the purpose of the meta-analysis, only studies using independent samples could be used. Some studies measured slightly different phenotypes in the same cohort, some authors examined more than one dependent measure concerning problematic cannabis use within one study, and some authors used a (sub)sample of the same cohort. In these cases only one of the reports was included in the meta-analysis, with a preference for reports with the largest sample, separate parameter estimates for each sex, the most suitable measure of cannabis use and estimates based on univariate models as opposed to multivariate models. Tables 1 and 2 show overviews of the available studies on cannabis use initiation and problematic use, respectively.

Finally, we used nine independent cohorts for males and eight for females for the meta-analysis of cannabis use initiation. For the meta-analysis of problematic cannabis use we used seven samples for males and six for females.

Meta-analysis

We meta-analysed the standardized variance components for the two phenotypes by calculating the weighted average genetic (A), shared environmental (C) and unshared environmental (E) estimates. An explanation of this method can be found in Li et al. [45] and Sutton et al. [46]. Briefly, to estimate the weighted mean, the male/female parameter estimates for each cohort were weighted by the number of males/females in the sample. In some cases the reports we used did not report separate parameter estimates for each sex (because they did not differ significantly). In these cases we used the equated estimates for both sexes. Calculations were conducted in Microsoft Office Excel 2007. Estimates were made separately for each sex and for all studies. In some cases we used the equated estimates for both sexes. Calculations were conducted in Microsoft Office Excel 2007. Estimates were made separately for each sex and phenotype (cannabis use initiation and problematic cannabis use). We also calculated the 95% CIs around each estimate, calculated from the variance in the sample of the source studies.

RESULTS

The twin studies we identified from the literature search, including information about the cohort, sample sizes, measure used and A, C and E estimates, are presented in Table 1 (cannabis use initiation) and Table 2 (problematic cannabis use). The studies selected for the meta-analyses are shown in bold type. One cohort [US, Vietnam Era Twin Registry (VETR)] used only male participants, so we could not include this cohort in our meta-analyses for females.

All cohorts are from western countries—more than half of them from US samples; other data were obtained in Australia, the United Kingdom, the Netherlands and Norway. Table 3 shows the results of the meta-analysis.

Initiation of cannabis use

For both sexes, individual differences in cannabis use initiation are due moderately to genetic, shared environmental as well as unshared environmental influences. Although the confidence intervals for male and female estimates overlap, additive genetic influences are somewhat stronger for males, while the shared environment plays a greater role in females. Figure 1a,b displays the results of the meta-analyses for genetic contributions to cannabis use initiation for males and females, respectively. The horizontal lines represent the 95% CIs around the heritability estimates (diamonds) from the different cohorts. When confidence intervals were not reported by the source studies, they were estimated (dotted lines), based on a logarithmic curve regression of the confidence intervals and sample size from the other studies. The bottom line shows the results of our meta-analysis, displaying the narrower confidence interval of the estimates compared to the intervals from the source studies. As can be seen, the point estimates from the meta-analyses fall within all confidence intervals from the source studies, suggesting that the source studies are homogeneous.

Problematic cannabis use

According to our meta-analysis, more than half the individual differences in problematic cannabis use are due to genetic variance, while shared environmental influences and unique environmental influences have substantially lower contributions. The A estimate is higher for females than for males but, again, confidence intervals overlap.

Compared to those on cannabis use initiation, genetic influences on problematic use are higher while shared environmental influences are lower for females. The most notable difference is the C effect for females, which explains only 15% of the variance for problematic use but almost 40% for initiation of cannabis use.

Figure 2a,b represents the meta-analyses for genetic contributions to problematic cannabis use for males and females, respectively. Again, the bottom line shows the results of our meta-analysis. The point estimates from the meta-analyses fall generally within the reported CIs of the estimates from the source studies, suggesting reasonable homogeneity of studies. Exceptions were Agrawal et al. [47], where the confidence intervals are particularly narrow, and McGue et al. [27] and Tsuang et al. [48], where confidence intervals were not reported but estimated by us.
Table 1 Overview of twin studies into life-time cannabis use. Studies in bold type are used in the present meta-analysis.

<table>
<thead>
<tr>
<th>Cohort</th>
<th>Reference</th>
<th>Country</th>
<th>Age (years)</th>
<th>Measure</th>
<th>Prevalence (in %)</th>
<th>Sex MZ (pairs)</th>
<th>DZ (pairs)</th>
<th>A (in %, ± 95% CIs)</th>
<th>C (in %, ± 95% CIs)</th>
<th>E (in %, ± 95% CIs)</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>Agrawal et al. 2008 [29]</td>
<td>Australia</td>
<td>24–36</td>
<td>Use prior to age 17</td>
<td>60.2 (life-time use)</td>
<td>M 494</td>
<td>395</td>
<td>72 (64–80)</td>
<td>–</td>
<td>28 (20–36)</td>
</tr>
<tr>
<td>1</td>
<td>Lynskey et al. 2007 [70]</td>
<td>Australia</td>
<td>24–36</td>
<td>Life-time use</td>
<td>Not reported</td>
<td>M/F 6265</td>
<td>individuals</td>
<td>60 (45–72)</td>
<td>9 (1–22)</td>
<td>32 (26–36)</td>
</tr>
<tr>
<td>1</td>
<td>Lynskey et al. 2002 [33]</td>
<td>Australia</td>
<td>24–36</td>
<td>Life-time use</td>
<td>M: 68.8</td>
<td>M 487</td>
<td>387</td>
<td>67 (9–75)</td>
<td>0 (0–50)</td>
<td>33 (25–43)</td>
</tr>
<tr>
<td>2</td>
<td>Agrawal et al. 2004 [71]</td>
<td>US, Virginia</td>
<td>M: 20–58</td>
<td>Early use, before age 18</td>
<td>52.9</td>
<td>M 1196</td>
<td>MM pairs</td>
<td>16 (52)</td>
<td>32 (17–26)</td>
<td>33 (25–43)</td>
</tr>
<tr>
<td>2</td>
<td>Karkowski et al. 2000 [73]</td>
<td>US, Virginia</td>
<td>M: 30.1 ± 7.6</td>
<td>Life-time use</td>
<td>Not reported</td>
<td>F 499</td>
<td>327</td>
<td>59 (44–73)</td>
<td>19 (9–31)</td>
<td>21 (16–32)</td>
</tr>
<tr>
<td>2</td>
<td>Kendler et al. 2003 [54]</td>
<td>US, Virginia</td>
<td>20–58</td>
<td>Life-time use</td>
<td>54.0</td>
<td>M 704</td>
<td>492</td>
<td>35 (35)</td>
<td>35 (35)</td>
<td>30 (35)</td>
</tr>
<tr>
<td>2</td>
<td>Kendler et al. 1999 [76]</td>
<td>US, Virginia</td>
<td>37.7 ± 7.5</td>
<td>Life-time use</td>
<td>48.7</td>
<td>F 1934</td>
<td>individuals</td>
<td>46 (18–77)</td>
<td>29 (9–48)</td>
<td>25 (20–32)</td>
</tr>
<tr>
<td>Cohort</td>
<td>Reference</td>
<td>Country</td>
<td>Age (years)</td>
<td>Measure</td>
<td>Prevalence (in %)</td>
<td>Sex</td>
<td>MZ (pairs)</td>
<td>DZ (pairs)</td>
<td>A (in %, ± 95% CIs)</td>
<td>C (in %, ± 95% CIs)</td>
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<tr>
<td>2</td>
<td>Neale et al. 2006 [78]</td>
<td>US, Virginia cohort 1</td>
<td>Not reported</td>
<td>Life-time use</td>
<td>Not reported</td>
<td>F</td>
<td>499</td>
<td>327</td>
<td>48 (0–78)</td>
<td>68 (15–93)</td>
</tr>
<tr>
<td>3</td>
<td>Maes et al. 1999 [28]</td>
<td>US, Virginia cohort 2</td>
<td>8–16</td>
<td>Life-time use</td>
<td>M: 0.0–12.6</td>
<td>M</td>
<td>300</td>
<td>184</td>
<td>22 (0–78)</td>
<td>68 (15–93)</td>
</tr>
<tr>
<td>4</td>
<td>Lessem et al. 2006 [79]</td>
<td>US, Add Health</td>
<td>Wave 1: M: 16.2 ± 1.7</td>
<td>Ever used</td>
<td>35</td>
<td>M/F</td>
<td>4846/4413</td>
<td>1092/945</td>
<td>21 (0–51)</td>
<td>57 (32–76)</td>
</tr>
<tr>
<td>5</td>
<td>Miles et al. 2001 [80]</td>
<td>US, Add Health</td>
<td>13–21</td>
<td>Ever used</td>
<td>35</td>
<td>F</td>
<td>144</td>
<td>131</td>
<td>31 (1–61)</td>
<td>47 (21–69)</td>
</tr>
<tr>
<td>6</td>
<td>McGue et al. 2000 [27]</td>
<td>US, Minnesota cohort 2</td>
<td>17</td>
<td>Ever used</td>
<td>35</td>
<td>F</td>
<td>159</td>
<td>133</td>
<td>39 (2–81)</td>
<td>44 (7–73)</td>
</tr>
<tr>
<td>7</td>
<td>Rhee et al. 2003 [34]</td>
<td>US, Colorado</td>
<td>12–19</td>
<td>Initiation</td>
<td>24</td>
<td>F</td>
<td>186</td>
<td>101</td>
<td>72 (29–95)</td>
<td>24 (2–65)</td>
</tr>
<tr>
<td>10</td>
<td>Grant et al. 2006 [68]</td>
<td>US, VETR</td>
<td>36–55</td>
<td>Ever used</td>
<td>~40</td>
<td>M</td>
<td>1583</td>
<td>1255</td>
<td>60 (8–82)</td>
<td>10 (0–28)</td>
</tr>
<tr>
<td>11</td>
<td>Tsuang et al. 1999 [82]</td>
<td>US, VETR</td>
<td>36–55</td>
<td>Transition from exposure to initiation</td>
<td>47.2 (life-time use)</td>
<td>M</td>
<td>1874</td>
<td>1498</td>
<td>44 (22–60)</td>
<td>10 (0–28)</td>
</tr>
<tr>
<td>12</td>
<td>Vink et al. (in press) [44]</td>
<td>the Netherlands</td>
<td>21–40</td>
<td>Cannabis initiation</td>
<td>36.2</td>
<td>M</td>
<td>158</td>
<td>98</td>
<td>44 (16–74)</td>
<td>31 (4–55)</td>
</tr>
</tbody>
</table>

A: additive genetic variance; C: shared environmental variance; E: unshared environmental variance; CIs: confidence intervals; DZ: dizygotic; F: female; M: male; MZ: monozygotic; OS: opposite sex; VETR: Vietnam Era Twin Registry.
Table 2 Overview of twin studies into problematic cannabis use. Studies in bold type are used in the present meta-analysis.

<table>
<thead>
<tr>
<th>Cohort</th>
<th>Reference</th>
<th>Country</th>
<th>Age (years)</th>
<th>Measure</th>
<th>Prevalence (%)</th>
<th>Sex</th>
<th>MZ (pairs)</th>
<th>DZ (pairs)</th>
<th>A (in %, ± 95% CIs)</th>
<th>C (in %, ± 95% CIs)</th>
<th>E (in %, ± 95% CIs)</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>Agrawal et al. 2007</td>
<td>Australia</td>
<td>24–36</td>
<td>Abuse/dependence</td>
<td>M: 28</td>
<td>M</td>
<td>487</td>
<td>387</td>
<td>68 (65–69)</td>
<td>14 (12–14)</td>
<td>18 (16–19)</td>
</tr>
<tr>
<td>2</td>
<td>Lynskey et al. 2002</td>
<td>Australia</td>
<td>24–36</td>
<td>Life-time abuse/dependence</td>
<td>M: 15.1</td>
<td>M</td>
<td>487</td>
<td>387</td>
<td>56 (23–77)</td>
<td>13 (0–40)</td>
<td>31 (21–45)</td>
</tr>
<tr>
<td>7</td>
<td>Kendler et al. 2007</td>
<td>US, Virginia</td>
<td>M: 36.3 (8.2) F: 37.0 (9.1)</td>
<td>Number of abuse/dependence symptoms</td>
<td>F: 9.4</td>
<td>F</td>
<td>1151 individuals</td>
<td>779</td>
<td>–</td>
<td>31</td>
<td></td>
</tr>
<tr>
<td>8</td>
<td>Kendler et al. 2003</td>
<td>US, Virginia</td>
<td>M: 23.7 ± 7.5</td>
<td>Life-time abuse/dependence</td>
<td>M: 16.6</td>
<td>M</td>
<td>704</td>
<td>492</td>
<td>73</td>
<td>1</td>
<td>26</td>
</tr>
<tr>
<td>11</td>
<td>Neale et al. 2006</td>
<td>US, Minnesota</td>
<td>Not reported</td>
<td>Abuse</td>
<td>F: 68</td>
<td>F</td>
<td>38</td>
<td>28</td>
<td>68</td>
<td>30</td>
<td>8</td>
</tr>
<tr>
<td>12</td>
<td>van den Bree et al. 1998</td>
<td>US, Minnesota</td>
<td>M: 28.5 F: 28.5</td>
<td>Abuse and or dependence</td>
<td>M: 24</td>
<td>F</td>
<td>56</td>
<td>66</td>
<td>24</td>
<td>9</td>
<td>47</td>
</tr>
<tr>
<td>Cohort</td>
<td>Reference</td>
<td>Country</td>
<td>Age (years)</td>
<td>Measure</td>
<td>Prevalence (%)</td>
<td>Sex</td>
<td>MZ (pairs)</td>
<td>DZ (pairs)</td>
<td>A (in %, ± 95% CIs)</td>
<td>C (in %, ± 95% CIs)</td>
<td>E (in %, ± 95% CIs)</td>
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<tr>
<td>5</td>
<td>Rhee et al. 2003 [34]</td>
<td>US, Colorado</td>
<td>12–19</td>
<td>One or more abuse/dependence symptoms</td>
<td>M: 13.7, F: 12.7</td>
<td>M</td>
<td>159</td>
<td>113</td>
<td>34 (0–67)</td>
<td>36 (10–60)</td>
<td>30 (16–48)</td>
</tr>
<tr>
<td>5</td>
<td>Young et al. 2006 [31]</td>
<td>US, Colorado</td>
<td>12–18</td>
<td>One or more abuse/dependence symptoms</td>
<td>M: 0.8–28.8, F: 0.0–22.9 (age 12–18)</td>
<td>M/F</td>
<td>645</td>
<td>702</td>
<td>55 (29–80)</td>
<td>24 (2–44)</td>
<td>21 (13–33)</td>
</tr>
<tr>
<td>6</td>
<td>Fu et al. 2002 [84]</td>
<td>US, VETR</td>
<td>33–52</td>
<td>Life-time dependence</td>
<td>6.6</td>
<td>M</td>
<td>1868</td>
<td>1492</td>
<td>50</td>
<td>13</td>
<td>37</td>
</tr>
<tr>
<td>6</td>
<td>True et al. 1999 [85]</td>
<td>US, VETR</td>
<td>33–52</td>
<td>3 or more life-time symptoms of abuse/dependence</td>
<td>6.7</td>
<td>M</td>
<td>1856</td>
<td>1479</td>
<td>43.9</td>
<td>21.3</td>
<td>35.8</td>
</tr>
<tr>
<td>6</td>
<td>Tsuang et al. 1999 [82]</td>
<td>US, VETR</td>
<td>36–55</td>
<td>Transition to abuse and dependence</td>
<td>7.2</td>
<td>M</td>
<td>1874</td>
<td>1498</td>
<td>22 (0–49)</td>
<td>9 (0–40)</td>
<td>69 (51–88)</td>
</tr>
<tr>
<td>6</td>
<td>Tsuang et al. 1998 [48]</td>
<td>US, VETR</td>
<td>36–55</td>
<td>Life-time abuse or dependence</td>
<td>7.2</td>
<td>M</td>
<td>1874</td>
<td>1498</td>
<td>33</td>
<td>29</td>
<td>38</td>
</tr>
<tr>
<td>6</td>
<td>Tsuang et al. 1996 [86]</td>
<td>US, VETR</td>
<td>36–55</td>
<td>Life-time abuse or dependence</td>
<td>7.2</td>
<td>M</td>
<td>1874</td>
<td>1498</td>
<td>33</td>
<td>29</td>
<td>38</td>
</tr>
<tr>
<td>7</td>
<td>Kendler et al. 2006 [30]</td>
<td>Norway</td>
<td>Mean = 28.2 (3.9)</td>
<td>Symptoms of abuse or dependence</td>
<td>1.6</td>
<td>M</td>
<td>220</td>
<td>117</td>
<td>75 (34–89)</td>
<td>0 (0–34)</td>
<td>23 (11–47)</td>
</tr>
</tbody>
</table>

A: additive genetic variance; C: shared environmental variance; E: unshared environmental variance; CIs: confidence intervals; DZ: dizygotic; F: female; M: male; MZ: monozygotic; OS: opposite sex; VETR: Vietnam Era Twin Registry.
DISCUSSION

Results of twin studies investigating the extent to which cannabis use vulnerability is due to genetic and environmental influences have been inconsistent. We carried out the first meta-analysis of twin studies into cannabis use in order to obtain more accurate estimates of the relative magnitude of genetic and environmental influences on cannabis use initiation and problematic use. Results for cannabis use initiation showed significant A, C and E influences accounting for 48%, 25% and 27% of the variance in males, and 40%, 39% and 21% of the variance in females. The corresponding A, C and E estimates for problematic cannabis use were 51%, 20% and 29% for males and 59%, 15% and 26% for females, all of which were significant. Confidence intervals for these estimates were considerably narrower than those in the source studies.

Our findings thus indicate that vulnerability to both cannabis use initiation and problematic use is substantially heritable. Twin studies that analysed both phenotypes in one model have revealed that part of these genetic factors overlap between cannabis use initiation and problematic use [49–51]. This implies that vulnerability to initiate cannabis use is due partly to the same set of genes as vulnerability to progress cannabis use.

For males, the relative genetic contribution was lower and the shared environmental contribution higher for initiation of cannabis use compared to problematic use, in accordance with Agrawal & Lynskey [43]. This may be because the initial stages of the process of cannabis use are more sensitive to environmental factors, such as drug availability and use by peers [52], whereas the likelihood of dependence is more influenced by biological factors such as individual differences in physical response to the drug.

Genetic factors influencing cannabis use overlap with those influencing use of other illicit drugs, although there are also specific genetic factors influencing use of each particular drug [31,48,53–55]. The general genetic vulnerability to drug use could be related to genes underlying personality characteristics such as novelty seeking [56–59], to biochemical attributes [60] or to psychiatric vulnerability [61,62].

By means of genetic linkage and association studies it could be possible to identify some of the specific genetic variants that influence cannabis use. However, cannabis use phenotypes are likely to be polygenic, with each gene accounting for only a small proportion of the variance, as seems to be the case for other complex phenotypes [63]. For substance use disorders in general, genome-wide association studies have found dozens of genes that could contribute to vulnerability [64]. Many of these gene variants are likely to alter specification and maintenance of neuronal connections [64]. Genes identified as affecting vulnerability to drug use problems could be potential targets for pharmaceutical drugs aiming to modify addictions.

The magnitude of C and E contributions to both cannabis use initiation and problematic use indicates that environmental factors (which are often modifiable) also play a substantial role. In a longitudinal twin study, Korthonen et al. [65] identified some of the environmental factors that predict cannabis as well as other illicit drug use. They found that paternal drinking behaviour was a significant familial predictor (although this could also be a manifestation of their shared genetic vulnerability). Other predictors they mention are early smoking onset, drinking to intoxication, having peers who smoke cigarettes or have acquaintances with drug experience and aggressive behaviours among males [65]. Scherrer et al. [66] also found that perceptions of substance use among siblings, friends and school peers are associated strongly with cannabis abuse/dependence in young adults.

Prevention and intervention programmes should focus upon identifying and modifying these risk factors. Thus, programmes focusing upon not just the individual, but also their family and peer groups, could be beneficial. Also, parents should be aware of the role they could play in preventing use of cannabis and other drugs by their children. Peers and parents can probably also serve as protective factors for cannabis use. Chabrol et al. [52] found that the number of peers opposed to cannabis use...
as well as students’ negative expectations of cannabis use were protective factors. Our finding that C influences seem to be more important to cannabis use initiation for females than for males suggests that females may be more sensitive to prevention and intervention programmes.

Although the parameter estimates from this meta-analysis have narrower confidence intervals than most of the source studies, their precision is still limited for several reasons. First, despite the large number of twin studies into cannabis use phenotypes, they were based only on a low number of independent cohorts; our variance components estimates are based on six to nine cohorts. Also, all cohorts are from western countries, with more than half of them from the United States, so the results are not necessarily generalizable to different populations. Additional and more varied cohorts would increase the

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**Figure 1** Heritability estimates [i.e. proportion of variance accounted for by genetic influences (A)] and 95% confidence intervals for the studies used in the meta-analysis of cannabis use initiation for males (a) and females (b). The bottom line shows the weighted A estimate and 95% confidence intervals estimated in the present meta-analysis. Dotted lines show confidence intervals estimated by a logarithmic curve regression on the sample sizes.
generalizability as well as the precision of the estimates. On the other hand, the countries from which the different samples are drawn have different policies towards cannabis use, which might influence the reported genetic and environmental estimates. For example, in countries with a more liberal cannabis policy, such as the Netherlands, cannabis is more easily available. Therefore, the relative contribution of environmental effects could be smaller and that of genetic effects larger. However, such a difference is not supported by the findings from Vink et al. [44], who reported A, C and E estimates for initiation for the Dutch sample that fit very well within the confidence intervals of our weighted averages.

Another limitation is that the twin studies we analysed differed regarding the composition of the sample, the phenotypic measures and the statistical method used. These inconsistencies between the studies are likely to be partly responsible for their inconsistent results. By combining the studies into one analysis we did not acknowledge possible differences between different samples and methods. For example, two source cohorts did not use population-based samples. First, in the meta-analysis for problematic use, we included a study using a treatment sample [67]. Because of the small sample size of this study and the fact that the reported variance components were relatively consistent with those from the other

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**Figure 2** Heritability estimates [i.e. proportion of variance accounted for by genetic influences (A)] and 95% confidence intervals for the studies used in the meta-analysis of problematic cannabis use for males (a) and females (b). The bottom line shows the weighted A estimate and 95% confidence intervals estimated in the present meta-analysis. Dotted lines show confidence intervals estimated by a logarithmic curve regression on the sample sizes.
source studies, this one study should not have biased our results strongly.

Secondly, for both variables we incorporated a male cohort of twins where both twins served in the US army during the Vietnam Era [48,68]. Genetic estimates based on this cohort are relatively high for cannabis initiation and relatively low for problematic cannabis use. Because the sample size of the source study is quite large, if the aetiology of cannabis use in Vietnam veterans is different from that in the general population, the inclusion of this study could have biased our results.

Also, in our analyses we combined studies using samples of varying age ranges, while results of earlier studies have suggested that from adolescence to adulthood the effect of shared environment gradually declines, while genetic influences gradually increases [69]. Younger individuals might have limited access to marijuana and experience less peer pressure. However, because of the low number of independent samples, it was not possible to distinguish meaningfully between different age groups or to conduct other subanalyses. The inclusion of adolescents and adults in one analysis might have modestly influenced our results, but estimates from the adolescent samples did not differ markedly from those from adult samples, so the effect would be small. Compared to the results of our analyses, heritability estimates for cannabis use initiation were relatively low and shared environmental estimates relatively high for some adolescent samples [27,28]. However, Rhee et al. [34] found a very high A estimate of 72%, and a relatively low C estimate for adolescent females.

Overall, our meta-analyses, by aggregating the results of a number of previous twin studies, provided more robust estimates of the genetic and environmental influences on cannabis initiation and problematic use. Because our analyses average estimates over samples of different sizes and demographic make-up, our findings are likely to be more generalizable than the source studies. Our results indicate that A, C and E factors each contribute significantly to vulnerability to both cannabis use initiation and problematic use. This confirms that cannabis problems do not have a single or simple cause, and suggests that both genetic and environmental factors are potential targets for treatment and prevention measures.

Declarations of interest
None.

Acknowledgements
K.J.H.V. is supported by DA018267 from the National Institute on Drug Abuse to M.T.L. J.M.V. is supported by NWO (VENI 451-06-004). S.E.M. is supported by an Australian NHMRC Sidney Sax Fellowship (443036). We thank Arpana Agrawal for advice regarding the Virginia twin sample, and Marianne van den Bree for providing additional results from their study not reported in the published text.

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